Statistical Machine Learning

https://cvml.ist.ac.at/courses/SML_W20

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Fall Semester 2020/2021 Lecture 2

Overview (tentative)

Date		no.	Topic
Oct 05	Mon	1	A Hands-On Introduction
Oct 07	Wed	2	Bayesian Decision Theory, Generative Probabilistic Models
Oct 12	Mon	3	Discriminative Probabilistic Models
Oct 14	Wed	4	Maximum Margin Classifiers, Generalized Linear Models
Oct 19	Mon	5	Estimators; Overfitting/Underfitting, Regularization, Model Selection
Oct 21	Wed	6	Bias/Fairness, Domain Adaptation
Oct 26	Mon	-	no lecture (public holiday)
Oct 28	Wed	7	Learning Theory I
Nov 02	Mon	8	Learning Theory II
Nov 04	Wed	9	Deep Learning I
Nov 09	Mon	10	Deep Learning II
Nov 11	Wed	11	Unsupervised Learning
Nov 16	Mon	12	project presentations
Nov 18	Wed	13	buffer

Refresher from Previous Lecture

The goal of (supervised) machine learning is

- use a training set $\mathcal{D} = \{(x^1, y^1), \dots, (x^n, y^n)\} \subset \mathcal{X} \times \mathcal{Y}$
- to find a prediction function $f: \mathcal{X} \to \mathcal{Y}$ (="learning")
- that works well on future data.

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Many learning techniques have been developed, such as

- decision trees
- (k-)nearest neighbor
- Perceptron
- Boosting \leftarrow today
- Artificial Neural Networks ← today

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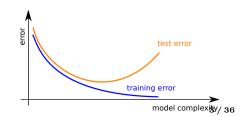
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Some phenomena are universal:

- models with too small complexity underfit the data
 - ▶ high training error, high test error
- models with too high complexity overfit the data
 - ▶ low training error, high test error



Given: training examples $\mathcal{D} = \{(x^1, y^1), \dots, (x^n, y^n)\} \subset \mathcal{X} \times \mathcal{Y}$ with $\mathcal{Y} = \{\pm 1\}$.

Problem:

- it's hard to guess a *strong* (=good) classifier.
- it's easy to guess weak (=slightly better than random) classifiers.

Question [Kearns, Valiant. 1988/89]:

• Given enough weak classifiers, can one always construct a strong one?

Answer [Schapire. 1990]:

Yes, by Boosting!

Boosting – Weak Classifiers

For example: if our features are

property	possible values
eye color	blue/brown/green
handsome	yes/no
height	short/tall
sex	male $(M)/female (F)$
soccer fan	yes/no

define (weak) classifiers:

$$h_1(x) = \begin{cases} +1 & \text{if eye color} = \text{brown} \\ -1 & \text{otherwise.} \end{cases} \qquad h_2(x) = \begin{cases} +1 & \text{if eye color} = \text{blue} \\ -1 & \text{otherwise.} \end{cases}$$

$$h_3(x) = \begin{cases} +1 & \text{if eye color} = \text{green} \\ -1 & \text{otherwise.} \end{cases} \qquad h_4(x) = \begin{cases} -1 & \text{if eye color} = \text{brown} \\ +1 & \text{otherwise.} \end{cases} , \ldots$$

$$h_5(x) = \begin{cases} +1 & \text{if handsome} = \text{yes} \\ -1 & \text{otherwise.} \end{cases} \qquad h_6(x) = \begin{cases} -1 & \text{if handsome} = \text{yes} \\ +1 & \text{otherwise.} \end{cases} , \ldots$$

Set of all possible combinations: $\mathcal{H} = \{h_1, \dots, h_J\}.$

AdaBoost - Training

input training set \mathcal{D} , set of weak classifiers \mathcal{H} , number of iterations T.

$$d_1 = d_2 = \cdots = d_n = 1/n$$
 (weight for each example) for t=1,...,T do

for $h \in \mathcal{H}$ do $e^t(h) = \sum\limits_{i=1}^n d_i \, \llbracket h(x^i) \neq y^i \rrbracket$ (weighted training error)

$$h_t = \mathbf{argmin}_{h \in \mathcal{H}} e^t(h)$$
 ("best" of the weak classifiers)

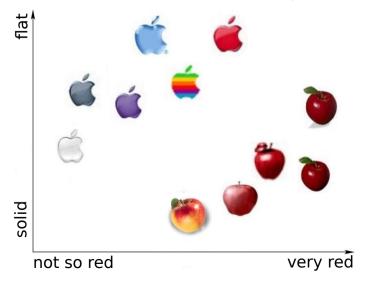
$$\alpha_t = \frac{1}{2} \log(\frac{1 - e_t(h_t)}{e_t(h_t)})$$
 (classifier importance, $\alpha_t = 0$ if $e_t(h_t) = \frac{1}{2}$)

for
$$i=1,\ldots,n$$
 do $\widetilde{d}_i \leftarrow d_i imes egin{cases} e^{lpha_t} & ext{if } h_t(x^i)
eq y^i, \\ e^{-lpha_t} & ext{otherwise}. \end{cases}$

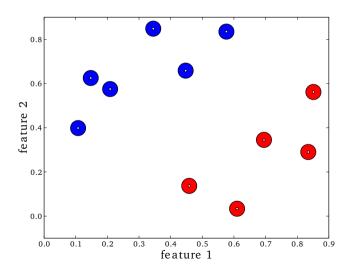
for
$$i=1,\ldots,n$$
 do $d_i \leftarrow \widetilde{d}_i/\sum_i \widetilde{d}_i$ end for

output classifier:
$$f(x) = \operatorname{sign} \sum_{t=1}^{T} \alpha_t h_t(x)$$

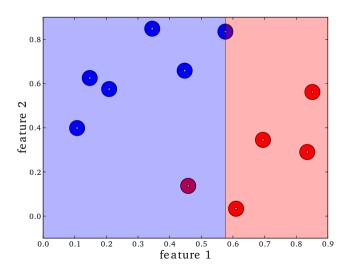
Task: $\mathcal{X} = \mathbb{R}^2$, weak classifiers look at each dimension separately



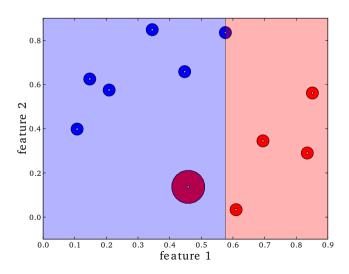
Iteration t=1, $d_1,\ldots,d_n=(\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11},\frac{1}{11})$



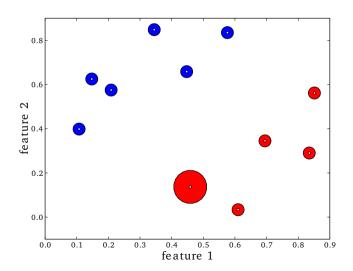
Iteration t=1, best weak classifier, $e_1(h_1)=\frac{1}{11}$, $\alpha_1=1.15$



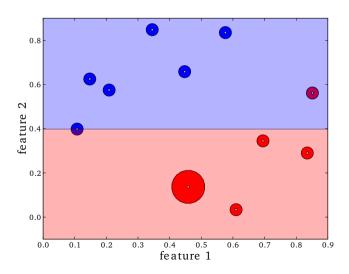
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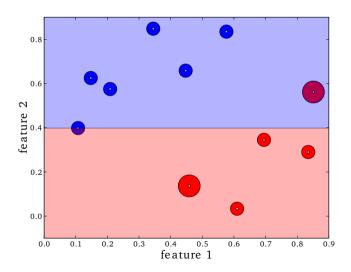
Iteration t=2, $d_1,\ldots,d_n\approx(\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{2},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20},\frac{1}{20})$



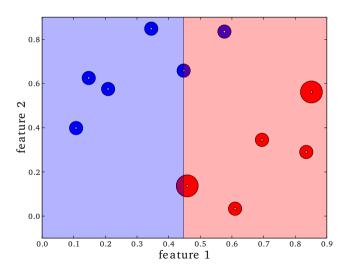
Iteration t=2, best weak classifier, $e_2(h_2)=\frac{1}{20}$, $\alpha_2=1.47$



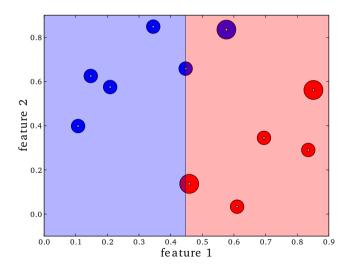
Iteration t=2, best weak classifier, $e_2(h_2)=\frac{1}{20}$, $\alpha_2=1.47$



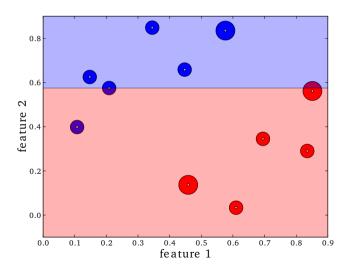
Iteration t = 3



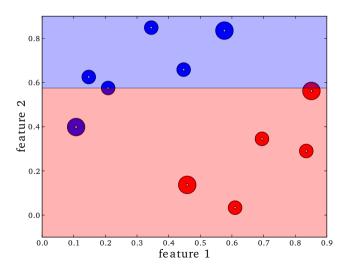
Iteration t = 3



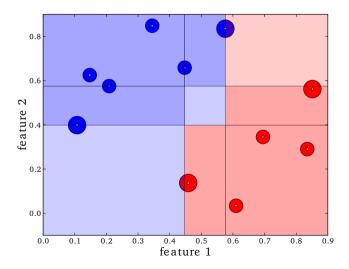
Iteration t=4



Iteration t=5



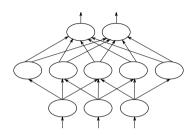
Final classifier: $f(x) = \text{sign} (1.15h_1(x) + 1.5h_2(x) + \dots + 0.9h_5(x))$



Artificial Neural Networks are predictive models inspired by (early) Neuroscience.

Main idea:

- stack layers of simple elements ("neurons")
- one layer's outputs are next layer's input.



Network parametrizes a function:

• each neuron N_i computes a linear/affine function

$$a_i = \langle w_i, x_{\mathsf{input}} \rangle + b_i$$

"activation"

followed by a componentwise non-linear transformation, $\sigma: \mathbb{R} \to \mathbb{R}$,

$$o_i = \sigma(a_i)$$
 e.g. $\sigma(t) = \max\{0, t\}$

Since approx. 2012, re-popularized under the name of $\textbf{Deep Learning} \rightarrow \text{lectures 7}$ and 8.

Understanding Machine Learning Methods

Decision Theory (for Supervised Learning Problems)

Goal:

- Understand existing algorithms
- Develop new algorithms with specific (optimal?) properties

For this, we'll rely on mathematics. Forget about implementation, finite data etc... (for now)

Notation

We treat all quantities of interest as random variables:

- input: random variable, X, taking values $x \in \mathcal{X}$ (we think of \mathcal{X} as continuous, but use discrete notation for simplicity)
- output: random variable, Y, taking values $y \in \mathcal{Y}$.
- joint probability distribution p(X = x, Y = y) = p(Y = y | X = x)p(X = x)
 - ▶ p(X = x): how likely is it that any $x \in \mathcal{X}$ will occur?
 - ightharpoonup p(Y=y|X=x): what's the probability that $y\in\mathcal{Y}$ is the correct answer for $x\in\mathcal{X}$?
- we write p(x,y) for of p(X=x,Y=y), p(y|x) instead of p(Y=y|X=x), etc.

Classification

First first look at classification, $\mathcal{Y} = \{1, \dots, M\}$, or $\mathcal{Y} = \{-1, +1\}$.

Question: What's the best classifier for a fully known problem?

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Definition (Generalization error)

Let $c: \mathcal{X} \to \mathcal{Y}$ be a decision rule. The *generalization error*, \mathcal{R} , of c is the probability of c making a wrong prediction, i.e.

$$\mathcal{R}(c) := \Pr_{(x,y) \sim p(x,y)} \{ c(x) \neq y \}.$$

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Definition (Bayes Classifier, Bayes Risk)

The smallest achievable generalization error,

$$\mathcal{R}_{\mathsf{Bayes}} = \min_{c:\mathcal{X} \to \mathcal{Y}} \mathcal{R}(c)$$

is called **Bayes** error. A classifier, c^* , that achieves the base error is called **Bayes classifier**.

Lemma

For any $x \in \mathcal{X}$ with p(x) > 0, a Bayes classifier has the decision rule

$$\hat{c}(x) \in \operatorname*{argmax}_{y \in \mathcal{Y}} p(y|x) \qquad \text{or (equivalently)} \qquad \hat{c}(x) \in \operatorname*{argmax}_{y \in \mathcal{Y}} p(x,y) \tag{*}$$

Proof. First: both rules are equivalent, because

$$\mathbf{argmax}_{y} \, p(x,y) = \mathbf{argmax}_{y} \, p(y|x) p(x) = \mathbf{argmax}_{y} \, p(y|x).$$

ullet 1) We rewrite the risk in terms of per-point contributions. For any $c:\mathcal{X} o \mathcal{Y}$

$$\mathcal{R}(c) = \Pr_{(x,y) \sim p(x,y)} \{ c(x) \neq y \}$$

$$= \underset{(x,y) \sim p}{\mathbb{E}} \llbracket c(x) \neq y \rrbracket$$

$$= \underset{x \sim p(x)}{\mathbb{E}} \underset{y \sim p(y|x)}{\mathbb{E}} \llbracket c(x) \neq y \rrbracket$$

$$= \underset{x \in \mathcal{X}}{\sum} p(x) \underbrace{\underset{y \sim p(y|x)}{\mathbb{E}} \llbracket c(x) \neq y \rrbracket}_{=:\mathcal{R}_{\sigma}(c)}$$

Lemma

For any $x \in \mathcal{X}$ with p(x) > 0, a Bayes classifier has the decision rule

$$\hat{c}(x) \in \mathop{\mathbf{argmax}} p(y|x)$$
 or (equivalently) $\hat{c}(x) \in \mathop{\mathbf{argmax}} p(x,y)$

• 2) For any $x \in \mathcal{X}$:

$$\mathcal{R}_{x}(\hat{c}) = \underset{y \sim p(y|x)}{\mathbb{E}} \llbracket \hat{c}(x) \neq y \rrbracket = \underset{y \sim p(y|x)}{\mathbb{E}} \llbracket \hat{c}(x) \neq y \rrbracket$$

$$= 1 - \underset{y \sim p(y|x)}{\mathbb{E}} \llbracket \hat{c}(x) = y \rrbracket$$

$$= 1 - \underbrace{p(Y = \hat{c}(x)|x)}_{\geq p(y|x) \text{ for all } y \in \mathcal{Y}}$$

$$\leq 1 - p(Y = c^{*}(x)|x)$$

$$= \underset{y \sim p(y|x)}{\mathbb{E}} \llbracket c^{*}(x) \neq y \rrbracket = \mathcal{R}_{x}(c^{*})$$

• 3) Let A be the set of points where c^* does not fulfill (*). Then, for all $x \in A$:

$$p(Y=\hat{c}(x)|x)>p(Y=c^*(x)|x)$$
 and $\mathcal{R}_x(\hat{c})<\mathcal{R}_x(c^*)$

(*)

Lemma

For any $x \in \mathcal{X}$ with p(x) > 0, a Bayes classifier has the decision rule

$$\hat{c}(x) \in \operatorname*{argmax}_{y \in \mathcal{Y}} p(y|x) \qquad \textit{or (equivalently)} \qquad \hat{c}(x) \in \operatorname*{argmax}_{y \in \mathcal{Y}} p(x,y) \qquad \qquad (*)$$

- 1) $\mathcal{R}(c) = \sum_{x \in \mathcal{X}} p(x) \mathcal{R}_x(c)$
- 2) for every $x \in \mathcal{X}$: $\mathcal{R}_x(\hat{c}) \leq \mathcal{R}_x(c^*)$
- 3) for every $x \in A$: $\mathcal{R}_x(\hat{c}) < \mathcal{R}_x(c^*) = \mathcal{R}_{\mathsf{Bayes}}$
- 4) Consequently, if there's at least one point in $x \in A$ with p(x) > 0, then

$$\mathcal{R}(\hat{c}) < \mathcal{R}(c^*)$$

But that inequality is impossible by the definition of $\mathcal{R}_{\mathsf{Bayes}}$. Therefore, no such point exists.

In summary: in all points x with p(x) > 0, the Bayes classifier fulfills (*).

In binary classification we can write c^* in closed form:

Lemma

For $\mathcal{Y} = \{-1, +1\}$, the Bayes classifier is given by

$$c^*(x) = \text{sign} \left[\log \frac{p(x, +1)}{p(x, -1)} \right],$$

as well as

$$c^*(x) = \text{sign} \left[\log \frac{p(+1|x)}{p(-1|x)} \right].$$

Proof: Exercise.

Should we use c^* to decide for every problem?

- ullet c^* is optimal when trying to *minimize the number of wrong decision*.
- That's often a good strategy, but not always.

Reminder

To evaluate a learning task, we use a loss function $\ell:\mathcal{Y}\times\mathcal{Y}\to\mathbb{R}.$

 $\ell(y,\bar{y})$ is the loss incurred when predicting \bar{y} if the correct answer is y.

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Reminder

To evaluate a learning task, we use a *loss function* $\ell: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$. $\ell(y, \bar{y})$ is the loss incurred when predicting \bar{y} if the correct answer is y.

Example: Doctor's dilemma

A patient has a cough but no fever. Should you make her a COVID19 suspect?

x: symptoms. $y \in \{\text{yes}, \text{no}\}$: COVID19

- $\ell(yes, yes) = 0$ (you did your job well)
- $\ell(no, no) = 0$ (you did your job well)
- $\ell(yes, no) = 50$ (the patent goes home and might infect many others)
- $\ell(no, yes) = 1$ (the patient has to take an unpleasant unnecessary test)

Common: one outcome is rare, but has bad consequences if mispredicted

Instead of minimizing the error probability, minimize the expected loss!

Definition

The classifier of minimal expected ℓ -risk is given by

$$c_{\ell}^{*}(x) := \mathbf{argmin}_{y \in \mathcal{Y}} \mathop{\mathbb{E}}_{\bar{y} \sim p(\bar{y}|x)} \ell(\,\bar{y},y\,).$$

Lemma

For
$$\mathcal{Y}=\{-1,+1\}$$
, and $\ell(y,\bar{y})$ given by the table $egin{array}{c|c} y\setminus\bar{y} & -1 & +1 \\ \hline -1 & a & b \\ \hline +1 & c & d \\ \hline \end{array}$

the risk w.r.t. ℓ is minimized by the decision rule

$$c_{\ell}^*(x) = \operatorname{sign}[-\log \frac{p(x,+1)}{p(x,-1)} + \log \frac{c-d}{b-a}-],$$
 or equivalently
$$c_{\ell}^*(x) = \operatorname{sign}[-\log \frac{p(+1|x)}{p(-1|x)} + \log \frac{c-d}{b-a}-].$$

Proof: Exercise...

Observation

The generalization error is the risk for 0/1-loss, i.e. $\ell(y, y') = [y \neq y']$.

Question: What's the best classifier for a fully known problem?

Question answered. We have identified the optimal classifiers!

Learning from Data

In the real world, p(x,y) is unknown, but we have a training set \mathcal{D} . What to do?

Definition

Given a training set \mathcal{D} , we call it

• a generative probabilistic approach: if we use \mathcal{D} to build a model $\hat{p}(x,y)$ of p(x,y), and then define

$$c(x) := \operatorname*{argmax}_{y \in \mathcal{Y}} \hat{p}(x,y) \quad \text{or} \quad c_{\ell}(x) := \operatorname*{argmin}_{y \in \mathcal{Y}} \underset{\bar{y} \sim \hat{p}(x,\bar{y})}{\mathbb{E}} \ell(\,\bar{y},y\,).$$

• a discriminative probabilistic approach: if we use $\mathcal D$ to build a model $\hat p(y|x)$ of p(y|x) and define

$$c(x) := \operatorname*{\mathbf{argmax}}_{y \in \mathcal{Y}} \hat{p}(y|x) \quad \text{or} \quad c_{\ell}(x) := \operatorname*{\mathbf{argmin}}_{y \in \mathcal{Y}} \underset{\bar{y} \sim \hat{p}(\bar{y}|x)}{\mathbb{E}} \ell(\bar{y}, y).$$

• a **decision theoretic approach**: if we use \mathcal{D} to directly seach for a classifier c.

Generative Probabilistic Models

Setting

We are given

• a training set of examples $\mathcal{D} = \{(x^1, y^1), \dots, (x^n, y^n)\}$, (note: technically rather a multi-set, elements can occur more than once)

Assumption:

• \mathcal{D} are independent and identically distributed (i.i.d.) samples from the unknown probability distribution p(x, y).

Shorthand notation,

- $m{\mathcal{D}}^X := \{x^1, \dots, x^n\}, \quad \text{input part of } \mathcal{D}$,
- $\mathcal{D}^Y := \{y^1, \dots, y^n\}$, output part of \mathcal{D} ,
- $\mathcal{D}_y := \{(x^i, y^i) \in \mathcal{D} : y^i = y\}$, all examples of label y.

Generative Probabilistic Models

Let's use $\mathcal D$ to form an estimate of p(x,y).

Definition

There's (at least) three approaches:

- parametric estimate:
 - fix a model class $\hat{p}(x, y; \theta)$,
 - estimate parameters $\hat{\theta}$ such that $\hat{p}(x,y;\hat{\theta}) \approx p(x,y)$.
 - lacktriangle the size of heta is independent of how large $\mathcal D$ is
- non-parametric estimate:
 - estimate any $\hat{p}(x,y) \approx p(x,y)$
 - lacktriangle the number of parameters/complexity of $\hat{p}(x,y)$ can grow with $|\mathcal{D}|$
- hybrids of the two

Generative Probabilistic Models: Multinomial

If $\mathcal X$ and $\mathcal Y$ are *finite*, we can represent any p(x,y) as a table of values.

To simplify notation, we look at generic $z \in \mathcal{Z}$ (think: z = (x, y)):

Definition (Empirical estimate)

Let z^1, \ldots, z^n be samples from p(z), then we call

$$\hat{p}_n(z) := \frac{1}{n} \sum_{i=1}^n [z^i = z]$$

the empirical estimate of p(z) from n samples.

Example: flipping a coin, $\mathcal{Z} = \{\text{heads}, \text{tails}\}$

- observed outcomes (n = 6): heads, heads, heads, tails, heads, tails
- $\hat{p}_6(\text{heads}) = \frac{1}{6}(1+1+1+0+1+0) = \frac{2}{3}$
- $\hat{p}_6(\mathtt{tails}) = \frac{1}{6}(0+0+0+1+0+1) = \frac{1}{3}$

Generative Probabilistic Models: Multinomial

Theorem (Convergence of the empirical estimate)

Let z^1, z^2, \ldots be i.i.d. samples from p(z). For every possible value $z \in \mathcal{Z}$

$$\Pr\left\{\lim_{n\to\infty}\hat{p}_n(z) = p(z)\right\} = 1.$$

Proof.

Every textbook on statistics: law of large numbers (strong version).

The curse of dimensionality

Setting:

Let $\mathcal{Z}=\mathcal{Z}_1\times\cdots\times\mathcal{Z}_d$, i.e. data decomposes into d non-trivial "features", "attributes", or "dimensions". Let $m_j:=|\mathcal{Z}_j|\geq 2$ for $j=1,\ldots,d$.

Lemma

The number of samples needed to estimate $\hat{p}(z)$ grows exponentially in d (unless we made additional assumptions).

Proof.

 $\hat{p}(z)$ has $|\mathcal{Z}| = \prod_{j=1}^d m_j \geq 2^d$ entries. Without further assumptions, each entry can be set arbitrarily, independently, except for the one constraint that they must sum to 1. Each sample influences only one bin, so we need at least $2^d - 1$ samples (in practice, many times that, of course).

Example (Dating agency table)											
	TRAINING	eyes	height	handsome	sex	soccer	date?				
	Apu	blue	tall	yes	male	no	yes				
	Bernice	brown	short	yes	female	no	no				
	:	÷	:	:	:	:	:				
	Itchy	brown	short	no	male	yes	yes				

Can we estimate p(x, y) here?

- $|\mathcal{X} \times \mathcal{Y}| = (3 \times 2 \times 2 \times 2 \times 2 \times 2) \times 2 = 96$, $\rightarrow p(x,y)$ has 95 free parameters
- We have 9 samples.
- Most possible combinations we have never seen!

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Bayes classifier from $\hat{p}(x,y)$: $c(x) := \mathbf{argmax}_{y \in \mathcal{Y}} \hat{p}(x,y)$

- $\hat{p}(\mathsf{Apu}, \mathtt{yes}) = \frac{1}{9}$, $\hat{p}(\mathsf{Apu}, \mathtt{no}) = 0$, o $c(\mathsf{Apu}) = \mathtt{yes}$,
- $\hat{p}(\mathsf{Jimbo}, \mathsf{yes}) = 0, \quad \hat{p}(\mathsf{Jimbo}, \mathsf{no}) = 0, \quad \rightarrow \quad c(\mathsf{Jimbo}) = ???,$

No clue about previously unseen patterns o very little generalization ability

Naive Bayes Model

Definition

Let $\mathcal{X} = \mathcal{X}_1 \times \cdots \times \mathcal{X}_d$. The Naive Bayes (NB) estimate of p(x,y) is

$$\hat{p}_{NB}(x,y) := \hat{p}(y) \prod_{j=1}^{d} \hat{p}_{j}(x_{j}|y),$$

where

- $\hat{p}(y)$ is an estimate of p(y),
- $\hat{p}_j(x_j|y)$ are estimates of $p(x_j|y)$ for every $j=1,\ldots,d$.

Naive Bayes Model

Definition

Let $\mathcal{X} = \mathcal{X}_1 \times \cdots \times \mathcal{X}_d$. The Naive Bayes (NB) estimate of p(x,y) is

$$\hat{p}_{NB}(x,y) := \hat{p}(y) \prod_{j=1}^{d} \hat{p}_{j}(x_{j}|y),$$

where

- $\hat{p}(y)$ is an estimate of p(y),
- $\hat{p}_j(x_j|y)$ are estimates of $p(x_j|y)$ for every $j=1,\ldots,d$.

Lemma

The number of free parameters in $p_{NB}(x,y)$ grows **linearly** with d (instead of exponentially).

Proof.

$$p_{NB}(x,y)$$
 has $|\mathcal{Y}|[1+\sum_{j=1}^d(m_j-1)]-1$ degrees of freedom.

Naive Bayes Classifier

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The Naive Bayes classifier is given by

$$c(x) := \underset{y \in \mathcal{Y}}{\operatorname{argmax}} \, \hat{p}_{NB}(x, y)$$

A Naive Bayes classifier needs much fewer examples for 'training' than one based on a full probability table.

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Remark

Even for $n \to \infty$, we likely won't have $\hat{p}_{NB}(x,y) \not\to p(x,y)!$

So, most likely, the NB model is wrong as a density estimate.

But that doesn't mean it doesn't work for making decisions!

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"All models are wrong, but some are useful." (George E. P. Box, 1979)

Parametric models for finite domains

Both models we saw so far are parametric:

For finite $z \in \mathcal{Z}$, p(z) is multinomial distribution:

- $|\mathcal{Z}|$ parameters: θ_z for $z \in \mathcal{Z}$ with $p(Z=z) = \theta_z$
- parameters fulfill
 - $\theta_z \ge 0$
 - $\triangleright \sum_{z} \theta_{z} = 1$

Similar for Naive Bayes model:

- $\hat{p}(y)$ is multinomial for $y \in \mathcal{Y}$, parameter $\theta_y \in \mathbb{R}^{|\mathcal{Y}|}$,
 - $\hat{p}(y) = \theta_y$ with $\theta_y \ge 0$, $\sum_{y \in \mathcal{V}} \theta_y = 1$,
 - $\hat{p}(x_j|y)$ is multinomial for $x_j \in \mathcal{X}_j$, parameters $\theta^j_{x_j}$
 - $\hat{p}(x_j|y) = \theta_{x_j}^y$ with $\theta_{x_j}^y \ge 0$, $\sum_{x_j \in \mathcal{X}_j} \theta_{x_j}^y = 1$, for all $y \in \mathcal{Y}$

We set parameters as $\theta_z = \frac{1}{n} \sum_{i=1}^n [z^i = z]$? Why?

Let $\hat{p}(z;\theta)$ be a parametric model with parameter $\theta \in \Theta$. Let $\mathcal{D}=\{z^1,\dots,z^n\}$ be i.i.d. samples from p(z).

Definition (Parameter estimation)

There's (at least) two main approaches to set θ :

Maximum Likelihood (ML) Estimation:

Which parameter value makes it most likely that we observed \mathcal{D} ?

$$\theta_{\mathsf{ML}} = \underset{\theta \in \Theta}{\mathbf{argmax}} \ \hat{p}(z^1, \dots, z^n; \theta) \ = \ \underset{\theta \in \Theta}{\mathbf{argmax}} \ \prod_i \hat{p}(z^i; \theta)$$

Maximum-A-Posteriori (MAP) Parameter Estimation:

Treat θ as a random variable itself. What's its most likely value given \mathcal{D} ?

$$\begin{split} \theta_{\mathsf{MAP}} &= \underset{\theta \in \Theta}{\mathbf{argmax}} \;\; p(\theta \mid z^1, \dots, z^n) \\ &= \underset{\theta \in \Theta}{\mathbf{argmax}} \;\; p(\theta) p(z^1, \dots, z^n | \, \theta) \; = \; \underset{\theta \in \Theta}{\mathbf{argmax}} \;\; p(\theta) \prod_i \hat{p}(z^i; \, \theta) \end{split}$$

where $p(\theta)$ is a *prior* distribution over the possible parameter values.

Parameter Estimation in Practice

Remark

In practice, one almost always uses the log-likelihood, which gives the same θ (because \log is a monotonous function):

$$\theta_{\mathsf{ML}} = \operatorname*{\mathbf{argmax}}_{\theta \in \Theta} \log \prod_{i=1}^{n} \hat{p}(x^{i}; \theta) = \operatorname*{\mathbf{argmax}}_{\theta \in \Theta} \sum_{i=1}^{n} \log \hat{p}(x^{i}; \theta)$$

and

$$\begin{split} \theta_{\mathsf{MAP}} &= \underset{\theta \in \Theta}{\mathbf{argmax}} && \log \ \left[\hat{p}(\theta) \prod_{i} \hat{p}(z^{i}; \, \theta) \right] \\ &= \underset{\theta \in \Theta}{\mathbf{argmax}} && \log \hat{p}(\theta) + \sum_{i} \log \hat{p}(z^{i}; \, \theta) \end{split}$$

Example in exercises: $z \in \{0,1\}$, $\hat{p}(z=1;\theta) = \theta$, $\hat{p}(z=0;\theta) = 1 - \theta$.